

Brownian motion

1 Introduction

Brownian motion is discussed in chapter 3 of [1].

2 Finite dimensional distributions of a stochastic process

Definition 2.1 (*Stochastic process*). A collection of random variables $\{\xi_t | t \geq 0\}$

$$\xi : \Omega \times \mathbb{R}_+ \rightarrow \mathbb{R}^d$$

is called a stochastic process.

Realizations of stochastic process are now paths rather than numbers:

Definition 2.2 (*Sample path*). For each $\omega \in \Omega$ the mapping

$$t \rightarrow \xi_t(\omega)$$

is called the sample path of the stochastic process.

In most applications stochastic processes are characterized by means of the family of all *finite dimensional joint distributions associated* to them. This means that for a stochastic process valued on \mathbb{R} , for any discrete sequence $\{t_i\}_{i=1}^n$ we consider the Borel σ -algebra $\mathcal{B}(\mathbb{R}^n)$ and $B_1, \dots, B_n \in \mathcal{B}$ and consider

$$P_{\xi_t}(B_1, t_1, \dots, B_n, t_n) \equiv P(\xi_{t_1} \in B_1, \dots, \xi_{t_n} \in B_n)$$

The so defined families of joint probability yield a consistent description of a stochastic process

$$\xi : \Omega \times \mathbb{R}_+ \rightarrow \mathbb{R}^d$$

if the following *Kolmogorov consistency conditions* are satisfied

- i $P(\mathbb{R}^d, t) = 1$ for any t
- ii $P_{\xi_t}(B_1, t_1, \dots, B_n, t_n) \geq 0$
- iii $P_{\xi_t}(B_1, t_1, \dots, B_n, t_n) = P_{\xi_t}(B_1, t_1, \dots, B_n, t_n, \mathbb{R}^d, t_{m+1})$
- iv $P_{\xi_t}(B_{\pi(1)}, t_{\pi(1)}, \dots, B_{\pi(n)}, t_{\pi(n)}) = P_{\xi_t}(B_1, t_1, \dots, B_n, t_n, \mathbb{R}^d, t)$

3 Wiener process

The above definitions allow us to characterize the Wiener process as a stochastic process:

Definition 3.1 (Wiener Process aka Brownian motion). A real valued stochastic process

$$w_t : \Omega \times \mathbb{R}_+ \rightarrow \mathbb{R}$$

is called a **Wiener process** or **Brownian motion** if

i $w_0 = 0$

ii any increment $w_t - w_s$ has Gaussian probability density

$$w_t - w_s \stackrel{d}{=} \frac{e^{-\frac{x^2}{2(t-s)}}}{\sqrt{2\pi(t-s)}} \quad (3.1)$$

for all $t \geq s \geq 0$.

iii For all times

$$t_1 < t_2 < \dots \leq t_n$$

the random variables

$$w_{t_1}, w_{t_2} - w_{t_1}, \dots, w_{t_n} - w_{t_{n-1}}$$

are independent (the process has independent increments).

3.1 Consequences of the definition

Some observations are in order

- It is not restrictive to consider the one dimensional case. A d -dimensional Wiener process a vector valued stochastic process whose components are each independent one-dimensional Wiener processes. More explicitly, the probability density of Brownian motion on \mathbb{R}^d is given by

$$p_{w_t}(\mathbf{x}) = \prod_{i=1}^d p_{w_t^i}(x_i)$$

- By i and ii we have that

$$p_{w_t}(x) = \frac{e^{-\frac{x^2}{2\sigma^2 t}}}{(2\pi\sigma^2 t)^{\frac{1}{2}}} \quad \& \quad p_{w_{t_2}-w_{t_1}}(x) = \frac{e^{-\frac{x^2}{2\sigma^2(t_2-t_1)}}}{[2\pi\sigma^2(t_2-t_1)]^{\frac{1}{2}}} \quad t_2 > t_1 \quad (3.2)$$

By iii The joint probability of w_{t_1} and $w_{t_2} - w_{t_1}$ is

$$p_{w_{t_1}, w_{t_2}-w_{t_1}}(x_1, y) = \frac{e^{-\frac{x_1^2}{2\sigma^2 t_1}}}{(2\pi\sigma^2 t_1)^{\frac{1}{2}}} \frac{e^{-\frac{y^2}{2\sigma^2(t_2-t_1)}}}{[2\pi\sigma^2(t_2-t_1)]^{\frac{1}{2}}}$$

By definition of probability density we can also write

$$p_{w_{t_1}, w_{t_2} - w_{t_1}}(x_1, y) = p_{w_{t_1}, w_{t_2} - x_1}(x_1, y) = p_{w_{t_1}, w_{t_2}}(x_1, y + x_1)$$

since

$$w_{t_2} = (w_{t_2} - w_{t_1}) + w_{t_1}$$

Recalling the definition of *conditional probability* we must also have

$$p_{w_{t_1}, w_{t_2}}(x_1, y + x_1) = p_{w_{t_2}|w_{t_1}}(x_1 + y, t_2 | x_1, t_1) p_{w_{t_1}}(x_1) \quad \forall x_1, x_2, t_2 > t_1$$

whence

$$p_{w_{t_2}|w_{t_1}}(x_1 + y, t_2 | x_1, t_1) = \frac{p_{w_{t_1}, w_{t_2}}(x_1, y + x_1)}{p_{w_{t_1}}(x_1)} = \frac{p_{w_{t_1}, w_{t_2} - w_{t_1}}(x_1, y)}{p_{w_{t_1}}(x_1)} = \frac{e^{-\frac{y^2}{2\sigma^2(t_2 - t_1)}}}{[2\pi\sigma^2(t_2 - t_1)]^{\frac{1}{2}}}$$

Finally, upon setting $x_2 = y + x_1$ we get into:

$$p_{w_{t_2}|w_{t_1}}(x_2, t_2 | x_1, t_1) = \frac{e^{-\frac{(x_2 - x_1)^2}{2\sigma^2(t_2 - t_1)}}}{[2\pi\sigma^2(t_2 - t_1)]^{\frac{1}{2}}}$$

3.2 Continuity and non-differentiability of the Wiener process

Proposition 3.1. • for each $\gamma \in (1/2, 1]$ and almost every realization ω (i.e. with $P = 1$)

$$w_t(\omega): \mathbb{R}_+ \mapsto \mathbb{R} \tag{3.3}$$

is nowhere Hölder continuous with exponent γ .

• The paths of Brownian motion are nowhere differentiable with probability one.

For each I For each I ; I and almost every , t $W(t,)$ is

Proof. Let us fix $N \in \mathbb{N}$ large enough that

$$n \left(\gamma - \frac{1}{2} \right) > 1 \tag{3.4}$$

We say that a path w_t is Hölder continuous with exponent γ for some $s \in I$ if

$$|w_t - w_s| \leq C |t - s|^\gamma \tag{3.5}$$

for all $t \in I$ and for some $C > 0$. Let us suppose $I = [0, 1]$. Let us now define the label

$$[N s] = \begin{cases} N s & \text{if } n s \in \mathbb{N} \\ i & \text{if } 1 > i - N s > 0 \end{cases} \tag{3.6}$$

for $N \in \mathbb{N}$ such that

$$N \gg n \tag{3.7}$$

The range of values

$$j = [Ns], \dots, [ns] \cdots + n - 1 \quad (3.8)$$

taken by the integer label j are then such that

$$s \leq \frac{j}{N} \leq s + \frac{n}{N} \leq 1 \quad (3.9)$$

We have then that if (3.5) holds true

$$\begin{aligned} \left| w_{\frac{j+1}{N}} - w_{\frac{j}{N}} \right| &\leq \left| w_{\frac{j+1}{N}} - s \right| + \left| w_{\frac{j}{N}} - s \right| \\ &\leq C \left| \frac{j+1 - Ns}{N} \right|^\gamma + C \left| \frac{j - Ns}{N} \right|^\gamma \leq \frac{2Cn^\gamma}{N^\gamma} \end{aligned} \quad (3.10)$$

We can therefore pick some integer M such that

$$2Cn^\gamma \leq M \quad (3.11)$$

for some integer M . Let us now define the event

$$F_{M,N}^i = \left\{ \omega \in \Omega \mid \left| w_{\frac{j+1}{N}} - w_{\frac{j}{N}} \right| \leq \frac{M}{N^\gamma}, \text{ for all } j = i, \dots, i+n-1 \right\} \quad (3.12)$$

for which (3.10) holds true. The Wiener process is Hölder continuous at *some time* $s \in [0, 1]$ if there exist an integer M and an integer N such that for any $k > N$ there exists some $i = 1, \dots, N$ such that

$$\mathbb{P}(F_{M,N}^i) > 0 \quad (3.13)$$

The formal expression of the aforementioned condition is

$$F = \bigcup_{M=1}^{\infty} \bigcup_{N=1}^{\infty} \bigcap_{k=N}^{\infty} \bigcup_{i=1}^N F_{M,N}^i \quad (3.14)$$

In particular if we show that

$$\bar{F}_{M,N} = \bigcap_{k=N}^{\infty} \bigcup_{i=1}^N F_{M,N}^i \quad (3.15)$$

has probability zero then we proved that the Wiener process is nowhere Hölder continuous with exponent γ . Since $\bar{F}_{M,N}$ is telescopic

$$\begin{aligned} \mathbb{P}(\bar{F}_{M,N}) &\leq \lim_{N \uparrow \infty} \inf_N \mathbb{P}\left(\bigcup_{i=1}^N \bar{F}_{M,N}^i\right) \\ &\leq \lim_{N \uparrow \infty} \sum_{i=1}^n \mathbb{P}(\bar{F}_{M,N}^i) \leq \lim_{N \uparrow \infty} N \left[\mathbb{P}\left(\left| w_{\frac{1}{N}} \right| \leq \frac{M}{N^\gamma}\right) \right]^n \end{aligned} \quad (3.16)$$

The last inequality follows because the Wiener product has independent and stationary increments

$$w_{\frac{j+1}{N}} - w_{\frac{j}{N}} \stackrel{d}{=} w_{\frac{1}{N}} \quad (3.17)$$

We can compute the probability

$$P\left(\left|w_{\frac{1}{n}}\right| \leq \frac{M}{N^\gamma}\right) = \int_{-\frac{M}{N^\gamma}}^{\frac{M}{N^\gamma}} dx \frac{e^{-\frac{Nx^2}{2}}}{\sqrt{2\pi N^{-1}}} \leq \tilde{K} N^{\frac{1}{2}-\gamma} \quad (3.18)$$

We see that it goes to zero for any $\gamma > 1/2$. Furthermore

$$P(\bar{F}_{M,N}) \leq \tilde{K} \lim_{N \uparrow \infty} N^{1+n(\frac{1}{2}-\gamma)} = 0 \quad (3.19)$$

since we hypothesized (3.4). □

3.2.1 Observations

- In probability Hölder continuity can be argued by elementary considerations

$$P(|w_t| > \epsilon) = 2 \int_{\epsilon}^{\infty} dx \frac{e^{-\frac{x^2}{2t}}}{\sqrt{2\pi t}} \leq 2 \int_{\epsilon}^{\infty} dx \frac{x}{\epsilon} \frac{e^{-\frac{x^2}{2t}}}{\sqrt{2\pi t}} = \sqrt{\frac{2t}{\pi}} \frac{e^{-\frac{\epsilon^2}{2t}}}{\epsilon} \quad (3.20)$$

For any

$$\epsilon = K t^\gamma \quad K > 0 \quad (3.21)$$

the probability tends to zero if $\gamma < 1/2$. The conclusion is that in probability

$$|w_t| \leq K t^\gamma \quad 0 < \gamma < \frac{1}{2} \quad (3.22)$$

- An indication of non-differentiability comes from the evaluation of the expected value of the absolute value of the incremental ratio of the Wiener process:

$$E\frac{|w_t|}{t} = 2 \int_0^{\infty} dx \frac{x}{t} \frac{e^{-\frac{x^2}{2t}}}{\sqrt{2\pi t}} = \sqrt{\frac{2}{\pi t}} \quad (3.23)$$

which diverges as t tends to zero.

- The probability density of

$$\eta_t = \frac{w_t}{t} \quad (3.24)$$

is

$$p(x) = \sqrt{\frac{t}{2\pi}} e^{-\frac{tx^2}{2}} \quad (3.25)$$

whence for any $0 \leq L < \infty$

$$\lim_{t \downarrow 0} P(-L \leq \eta \leq L) = \lim_{t \downarrow 0} \int_{-L}^L dx \sqrt{\frac{t}{2\pi}} e^{-\frac{tx^2}{2}} = 0 \quad (3.26)$$

This fact indicates that the probability concentrates at infinity.

4 Markov processes and Chapman-Kolmogorov equation

A special important class of non-anticipating stochastic processes are Markov processes

Definition 4.1. Let $\mathcal{F}_{[0,t]}^\xi$ the filtration generated by a stochastic process ξ_t . Then ξ_t is Markov if for any $t \geq s$ and any event A

$$P(\xi_t \in A | \mathcal{F}_{[0,s]}^\xi) = P(\xi_t \in A | \mathcal{F}_s^\xi) = P(\xi_t \in A | \xi_s) \quad (4.1)$$

This means that the state of the system at time s fully specify further evolution independently of what happened before for times smaller s . The system has no “memory”. In particular if the transition probability p density of the Markov process $\xi_t : \Omega \times I \mapsto \mathbb{R}$ is available

$$P(\xi_t \in A | \xi_s = y) = \int_A dx p(x, t | y, s) \quad (4.2)$$

then the definition of Markov process implies that the Chapman-Kolmogorov equation

$$p(x_2, t_2 | x_0, t_0) = \int_{\mathbb{R}} dx_1 p(x_2, t_2 | x_1, t_1) p(x_1, t_1 | x_0, t_0) \quad (4.3)$$

must hold true for any (x_2, x_0) and for any $t_i, i = 0, 1, 2$.

4.1 Quadratic variation of the Wiener process

In the case of the Wiener process we have

Proposition 4.1 (*Quadratic variation of the B.M.*). The quadratic variation of the Brownian motion in $[0, t]$ for any $t \in \mathbb{R}_+$

$$V_{[w,w]}(I) = t$$

in the sense of $\mathbb{L}^2(\Omega)$.

Proof. By direct calculation we know that

$$\mathbb{E} w_t^2 = t \quad (4.4)$$

Let \mathfrak{p}_n a partition paving $[0, t]$ with n sub-intervals:

$$Q_n := \sum_{\mathfrak{p}_n} (w_{t_k} - w_{t_{k-1}})^2$$

we have then

$$\mathbb{E} (Q_n - t)^2 = \mathbb{E} \sum_{kl \in \mathfrak{p}_n} [(w_{t_k} - w_{t_{k-1}})^2 - (t_k - t_{k-1})] [(w_{t_l} - w_{t_{l-1}})^2 - (t_l - t_{l-1})]$$

For non-overlapping intervals, the averaged quantities are independent random variables with zero average. The only contributions to the sum come from overlapping intervals:

$$\mathbb{E} (Q_n - t)^2 = \mathbb{E} \sum_{k \in \mathfrak{p}_n} [(w_{t_k} - w_{t_{k-1}})^2 - (t_k - t_{k-1})]^2 = 2 \sum_{k \in \mathfrak{p}_n} (t_k - t_{k-1})^2$$

whence

$$\mathbb{E} (Q_n - t)^2 \leq 2t \max_{k \in \mathfrak{p}_n} (t_k - t_{k-1}) \xrightarrow{\max_{k \in \mathfrak{p}_n} (t_k - t_{k-1}) \downarrow 0} 0$$

□

The finite value of the quadratic variation motivates the estimate

$$dw_t \sim O(\sqrt{dt})$$

for typical increments of the Wiener process.

References

- [1] L. C. Evans. An introduction to stochastic differential equations. UC Berkeley, Lecture Notes, 2006.